

Seasonal effects in the long-term correction of short-term wind measurements using reanalysis data

Alexander Basse^{1,2}, Doron Callies^{1,2}, Anselm Grötzner³, and Lukas Pauscher²

¹Department of Integrated Energy Systems, University of Kassel, Wilhelmshöher Allee 73, 34121 Kassel, Germany

²Fraunhofer Institute for Energy Economics and Energy System Technology (IEE), Königstor 59, 34119 Kassel, Germany

³Ramboll Deutschland GmbH, Elisabeth-Consbruch-Straße 3, 34131 Kassel, Germany

Correspondence: Alexander Basse (alexander.basse@uni-kassel.de)

Abstract.

Measure-Correlate-Predict (MCP) approaches are often used to correct wind measurements to the long-term wind conditions on site. This paper investigates systematic errors in MCP-based long-term corrections which occur if the measurement on site covers only a few months (seasonal biases). In this context, two common linear MCP methods are tested and compared, namely

5 Variance Ratio and Linear Regression with Residuals. Wind measurement data from 18 sites with different terrain complexity in Germany are used (measurement heights between 100 and 140 m). Six different reanalysis data sets serve as the reference (long-term) wind data in the MCP calculations. Besides experimental results, theoretical considerations are presented which provide the mathematical background for understanding the observations. General relationships are derived which can be used to trace the seasonal biases to the mechanics of the methods and the properties of the reanalysis data sets. This allows the

10 transfer of the results of this study to different measurement durations, other reference data sets and other regions of the world. In this context, it is shown both theoretically and experimentally that the results do not only depend on the selected reference data set but also significantly change with the choice of the MCP method.

1 Introduction

An extensive measurement campaign generally constitutes an essential part of wind resource assessment and, therefore, of a

15 successful wind energy project. In most cases, these measurements provide around one year of wind data at the site of interest (Lackner et al., 2008). Inter-annual variations in wind speed are reported to vary by between 4 % and up to 10 % (e.g., Corotis, 1976; Justus et al., 1979; Klink, 2002), depending on the respective site; hence, the measured wind data usually do not represent the long-term wind conditions. This aspect becomes even more momentous when the energy in the wind is considered which has been reported to vary by 6 % (Pryor et al., 2018) up to 20 or even 30 % (Corotis, 1976; Albrecht and Klesitz, 2006; Pryor

20 et al., 2006) from year to year. To account for this issue, a long-term correction is performed.

For this purpose, reference data are needed which should be available for a long-term period of one to two decades (Lackner et al., 2008; Carta et al., 2013; Liléo et al., 2013) and show a high degree of similarity to the measured wind data (e.g., a high correlation coefficient of measured and reference data).

Over the recent past, reanalysis data gained more and more popularity in the wind industry and are now used extensively
25 in wind resource assessment (Miguel et al., 2019; Ramon et al., 2019). Reanalysis data sets are produced using numerical
weather simulations with a fixed state-of-the art model and assimilating historical weather data. In contrast to models used for
weather prediction, which are often updated and changed during operations, they therefore provide temporally consistent data
sets over periods of up to several decades. Different types of reanalysis data are available, ranging from (often freely available)
global data sets (e.g., MERRA-2 by NASA (NASA, 2019), ERA5 by ECMWF (CDS, 2018)) to mesoscale reanalyses, which
30 are generally not free of charge but provide higher spatial resolution.

A statistical procedure relating the reference data to the measured data is performed to derive a correction function. In this
context Measure-Correlate-Predict (MCP) approaches have evolved to become a standard tool for wind farm developers (Carta
et al., 2013). These methods model a statistical relationship between the time series of the reference and the measurement
data. Afterwards, the relationship is applied to the long-term reference data, providing the long-term wind conditions. The
35 relationship between reference and target data, therefore, is assumed not to be time-dependent, i.e., valid in the correlation
period as well as in the correction period.

Numerous MCP methods are used in modern wind resource assessment applications. They range from simple linear models
(e.g., García-Rojo, 2004; Rogers et al., 2005a; Romo Perea et al., 2011; Weekes and Tomlin, 2014a) to complex machine
learning approaches like neural networks (e.g., Bass et al., 2000; Albrecht and Klesitz, 2006; Bilgili et al., 2007; Velázquez
40 et al., 2011; Jie Zhang et al., 2014). The investigation and comparison of different MCP approaches has been subject to a large
amount of scientific publications. In Carta et al. (2013) an extensive review is given on existing MCP methods applied in wind
resource assessment and related research fields. It is concluded that, by far, the most commonly used MCP methods in the wind
industry are based on linear approaches. Other studies confirm this observation and underline the benefit of the simplicity of
linear MCP methods for use in wind energy applications (e.g., Sørensen et al., 2011; Weekes and Tomlin, 2014c; Weekes et al.,
45 2015). In a round-robin experiment in Germany in 2018 it was found that 24 of 29 consultants used linear correlation methods,
which mostly outperformed more complicated approaches (Basse et al., 2018).

In order to enable a precise determination of the relationship between measurement and reference data, a sufficient amount of
measurement data is necessary, that is, the concurrent period needs to be "long enough". Various studies have been presented in
which the question is addressed of how long the time span covered by the measurement should be. In general, it is recommended
50 to be at least one year (Carta et al., 2013), while the use of complete years is important as an uneven representation of different
months increases the uncertainty (Taylor et al., 2004; Liléo et al., 2013). As a consequence of such studies, an amount of 12
months of measurement is recommended or even a mandatory minimum duration due to technical guidelines and standards
such as FGW e.V. (2020), IEC (2017) or MEASNET (2016).

From an economic perspective, though, there is a strong desire to reduce the duration of the measurement in order to save
55 time and money (Carta et al., 2013). This is especially true with the increasing popularity of lidar measurements, which have
a high mobility and low installation costs compared to classical measurement masts while comparatively high running costs.
Moreover, an estimate of the wind conditions on site is often of interest for the wind park planner before the measurement

campaign is completed. In all such cases, a smaller amount of wind data needs to be dealt with and a long-term correction is performed based on wind measurement data which comprise much less than a year.

- 60 However, seasonal effects occur when the measurement does not cover all seasons (Rogers et al., 2005a; Saarnak et al., 2014; Weekes and Tomlin, 2014a,b,c) resulting in a dependence of the estimated energy yield on the period in which the measurement is conducted. These can induce systematic deviations and, thus, increase the uncertainty of the resource assessment significantly. Therefore, understanding seasonal patterns in long-term correction and their relation to data sources and the choice of the MCP method is of high interest for the wind industry.
- 65 Several studies have investigated the accuracy of a long-term correction (LTC) of short-term wind measurements in dependence of the measurement duration (e.g., Taylor et al., 2004; Rogers et al., 2005a,b; Romo Perea et al., 2011; Weekes and Tomlin, 2014c; Weekes et al., 2015; Miguel et al., 2019). While in some of these, seasonal effects are broadly addressed, to the authors' knowledge there is a lack of scientific publications which give profound explanations for seasonal patterns in biases of the LTC. This paper investigates seasonal effects and related biases in wind speed (mean and variance) and annual energy yield
- 70 in the LTC induced by short (three months) measurement periods. Motivated by their relevance for practical use, two linear MCP methods are applied and compared: Linear Regression with Residuals (Weekes and Tomlin, 2014a) and the Variance Ratio method (Rogers et al., 2005a). First, theoretical considerations are developed to assess the impact of varying statistical relationships between the measurement and the reference data in the short-term period when compared to the long-term period. In a second step, wind measurement data from 18 sites in Germany and six different reanalysis data sets are used to assess the
- 75 significance and magnitude of seasonal effects in the LTC. Interrelations of the seasonal effects with properties of the reference data and the correlation method are analyzed both theoretically and experimentally.

2 Measurement and reanalysis data used in this study

- An overview of the measurement campaigns is given in Tab. 1. All sites are located in Germany; the complexity of the sites ranges from flat agricultural areas to the hilly low mountain ranges in Central Germany (one of the complex sites is described
- 80 in Pauscher et al. (2018)). For all sites a time series of an entire year for a height level between 100 and 140 m is available, representing typical hub heights of modern wind turbines. The data were collected by profiling lidar (Light Detection And Ranging, see e.g., Emeis et al. (2007)) of type Leosphere WindCube V1 & V2 (Leleu, 2019), sodar (Sound Detection And Ranging, see e.g., Bradley (2008)) or mast measurements. The one-year periods are distributed relatively homogeneously between May 2013 and April 2019; only the year 2016 may be judged slightly over-representated (with eight of the 18 sites
- 85 covering at least a few months of the year 2016). The measurement data were collected at a temporal resolution of 10 minutes and then averaged to hourly values (centered at the full hour) to comply with the typical temporal resolution of the reanalysis data (see below). The availability of the measurement data is higher than 80 % at all sites with more than 90 % data availability at 14 sites. All data gaps are smaller than 100 consecutive hours except for a single site (Site 17 in Tab. 1), where approx. 10 days of data are missing in winter (overall data availability for this site: 95 %).

Table 1. Details of the measurement sites. The duration of the individual measurements is exactly one year. The measurements were carried out between May 2013 and April 2019.

Site No.	orography and surface cover	measurement height [m]	measurement device
1	hilly, forested	140	Lidar (WindCube V2)
2	slightly hilly, forested	140	Lidar (WindCube V2)
3	mainly flat, forested	140	Lidar (WindCube V2)
4	hilly, sparsely forested	140	Lidar (WindCube V1)
5	slightly hilly, barely forested	140	Lidar (WindCube V1)
6	slightly hilly, forested	140	Lidar (WindCube V2)
7	hilly, forested	140	Lidar (WindCube V1)
8	slightly hilly, no trees	140	Lidar (WindCube V1)
9	slightly hilly, sparsely forested	140	Lidar (WindCube V1)
10	mainly flat, buildings nearby	135	Lidar (WindCube V2)
11	mainly flat, small town nearby	140	Lidar (WindCube V2)
12	hilly, forested	135	Mast
13	slightly hilly, forested	140	Mast
14	rather flat, forested	130	Mast
15	flat, close to a city	110	Mast
16	flat, agricultural area	100	Mast
17	rather flat, forest nearby	140	Sodar
18	slightly hilly, forested	140	Sodar

90 The following six different reanalysis data sets serve as reference data in the MCP calculations:

1. **MERRA-2** (GMAO, 2015). The Modern-Era Retrospective Analysis for Research and Applications Version 2 (MERRA-2) is based on global numerical weather analyzes of the U.S. National Aeronautics and Space Agency (NASA). The data are available as one-hour time series since 1980 for a height of 50 m and a spatial resolution of $0.5^\circ \times 0.66^\circ$. The time stamps refer to average hourly values centered at 00:30 h, 01:30 h etc. In order to obtain comparability with the other
95 reanalysis data sets and consistency in temporal terms, these were interpolated to values centered at the full hour.
2. **ERA5** (Hersbach et al., 2020). The data set is calculated at the European Centre for Medium-Range Weather Forecasts (ECMWF) and provided by the Copernicus Climate Change Service. The ERA5 data represent the follow-up data set to the ERA-Interim reanalyses of the ECMWF. The spatial resolution of the ERA5 data is approx. 31 km ($\approx 0.28^\circ$). Long-term series of this data set are available for 100 m above ground in an hourly resolution. In contrast to the MERRA-2
100 data, these data are instantaneous values instead of averaged wind speeds (centered at the full hour).
3. **EMD-ConWx** (EMD, 2020a). This data set is created using the WRF model (Weather Research & Forecasting Model, see WRF (2020)) and is provided by EMD International A/S from Denmark. It is based on the ERA-Interim reanal-

ysis data of the ECMWF, refined to a resolution of 3 km. The temporal resolution of the long-term time series is 1 h (instantaneous values centered at the full hour). Wind data are provided at heights of 10, 25, 50, 75, 100, 150, and 200 m.

- 105 4. **EMD-WRF Europe+** (EMD, 2020b). This dataset is a further development of the EMD-ConWx data. The ERA5 reanalysis data have replaced the ERA-Interim data, while spatial resolution and temporal properties have not changed. Wind data are provided at the same heights as in EMD-ConWx and six additional heights up to 4000 m.
- 110 5. **anemosM2**: anemos Windatlas based on MERRA-2 (anemos, 2020a,c). Similar to the EMD data sets, these data are created based on a downscaling of global reanalysis data (here: MERRA-2) using the WRF model (version 3.7.1) to a resolution of 3 km. In contrast to the other models, anemos uses statistical post-processing based on measurement data, known as remodeling, to improve the simulation results. Furthermore, additional downscaling of the data from the 3 km grid to the specific site is applied. The heights of the wind data are generally freely selectable between 40 and 200 m; for the analysis in this study, wind data at 100 and 140 m were provided.
- 115 6. **anemosE5**: anemos Windatlas based on ERA5 (anemos, 2020b,c). This data set is similar to the anemosM2 but uses ERA5 data. Furthermore, in the course of the remodeling, a seasonal correction is performed, i.e., biases in the annual cycle of the ERA5 data are corrected before the statistical downscaling is implemented. The goal is to better capture the seasonal behaviour of the wind conditions. Additionally, a more precise consideration of the roughness at the respective site represents a further difference to the anemosM2 data. Both the magnitude of the seasonal corrections as well as the modifications on roughness constitute a trade secret of anemos (anemos, 2021).

120 It should be noted that both the anemosM2 and anemosE5 models generally provide a temporal resolution of 10 minutes. In order to guarantee comparability of the results, these were averaged to 1 h ensuring the same temporal resolution for all reanalysis data sets.

In general, reanalysis data are modeled for different locations on a geographical grid. In this study, data were selected from the grid point closest to the respective site. For data sets 3. - 6. data at more than one height level were provided. In these cases, 125 the data at the height closest to the measurement were used (i.e., 100 and 150 m for EMD-ConWx and Emd-Wrf Europe+, 100 and 140 m for the two anemos data sets). For the MERRA-2 and ERA5 data sets the data at the given height (i.e., 50 and 100 m, respectively) were used, i.e., no vertical extrapolation (or interpolation) was performed in this study.

3 Methodology

This study compares statistics as observed over different periods in the investigated data - namely short-term data and long-term data. For this purpose, the convention is applied that capital letters are used for long-term variables (e.g., the long-term 130 corrected wind speed) while parameters in lower case letters represent data from the short-term period. The indices "meas", "ref", and "corr" refer to measurement, reference (i.e., reanalysis) and corrected data, respectively.

3.1 Selection of short-term periods and procedure of long-term correction

Short-term periods with a duration of 90 consecutive days are selected starting at the first day of year and running through the data with an increment of three days ("sliding window", e.g., the first period starts on January 1, the second on January 4 etc.). When the end of the data is reached, the data from the beginning of the data set is appended ensuring that 122 90-day measurement periods can be investigated for all reanalysis data sets and all sites.

In a first step, the data in these three-month data portions are analyzed with respect to, e.g., mean and variance of wind speed (Sect. 5.1, 5.2 and 5.3). Secondly, MCP predictions are performed. Regression parameters are derived using the short-term data and, afterwards, correction is performed in the entire one-year period in which measurement data are available. Finally, the corrected data are compared to the measured one-year data (benchmark) and error scores are derived (see Sect. 3.3). The general procedure is illustrated in Fig. 1. The results, therefore, do not represent the overall errors (or uncertainty) of an LTC in general, which is usually performed over a period of ten years or more (Lackner et al., 2008; Carta et al., 2013; Liléo et al., 2013). Instead, the analysis provides findings on systematic errors (seasonal biases) which emerge due to the reduction of the measurement duration from one year to three months.

It should be noted that in practical applications, a sector-wise regression is often performed for an LTC of measurement data comprising a whole year. This means, that the regression parameters are calculated separately for different wind direction bins which allows to take the effects of terrain on wind flow into account. This can be important especially in a complex environment (López et al., 2008). For the shorter three-month periods, sectorwise binning, however, generally yielded slightly worse results in this study (presumably due to low data coverage in the different direction sectors). This procedure is, therefore, not applied here. It is acknowledged, though, that in some specific cases a sectorwise approach can be a reasonable choice for an LTC of short-term measurements nevertheless.

When a correction is performed, few negative wind speed values can occur. In this study, these values were set to zero. In order to derive robust, conclusive findings, the individual results obtained at the 18 sites were averaged arithmetically, resulting in one set of statistics (e.g., error scores) for each reanalysis data set and each 90-day measurement period.

As mentioned in the introduction, the correlation coefficient of site and reference data should be evaluated before a long-term correction is performed. It is obvious that the correlation coefficient is lower when considering short-term periods (this will shortly be addressed in Sect. 5.4.2). In most combinations of reanalysis and site data, the correlation coefficient was $r_{\text{ref, meas}} > 0.65$ throughout, despite the small amount of only 90 days of data. Only in case of the EMD-ConWx and EMDWrf Europe+ datasets, values of less than 0.5 were observed in summer periods at some sites. This should be considered when assessing the results. However, it should be noted that this work intends to analyze the effects of shortening the measurement campaign for MCP approaches. Therefore, periods with low correlation coefficients are not excluded but the effects of the correlation coefficient are explored in several sections (Sect. 4.3, 5.2 and 5.4.2 in particular).

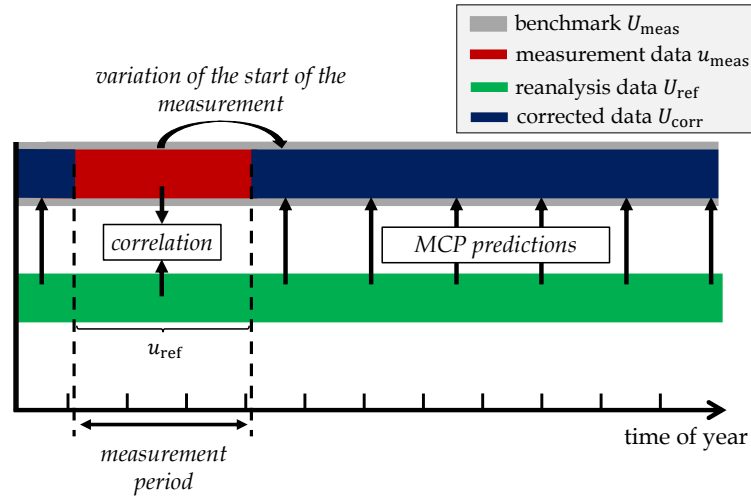


Figure 1. Illustration of the general procedure used in this study regarding the MCP predictions. In the short-term (measurement) period, a correlation function of the measured (u_{meas}) and reanalysis data (u_{ref}) is determined. This relationship is used to correct the reanalysis data in the entire one-year period, U_{ref} . Finally, the obtained corrected data U_{corr} is compared to the actually measured values (benchmark) U_{meas} in order to estimate the accuracy.

3.2 Long-term correction: Measure-Correlate-Predict (MCP) approaches

165 In this section, a brief overview of the two MCP methods used in this study is given. Both implement a linear model to derive a relation between measurement (u_{meas}) and reference wind speed (u_{ref}) in the measurement period. This linear relationship is generally expressed in the form

$$u_{\text{meas}} = \beta_0 + \beta_1 \cdot u_{\text{ref}} + \varepsilon, \quad (1)$$

170 where β_0 and β_1 represent the main regression parameters. ε indicates the residuals (deviations from data points to fitting line, see e.g., Ellison et al. (2009)).

3.2.1 Linear regression with residuals

The probably most widely used linear model is simple linear regression. In this approach the respective regression parameters $\beta_{0,\text{LR}}$ and $\beta_{1,\text{LR}}$ are calculated via the linear least squares method which minimizes the average squared deviation of the data points from the fitting line (see e.g., Draper and Smith, 1998). This results in

$$175 \quad \beta_{1,\text{LR}} = r_{\text{ref,meas}} \cdot \frac{\sigma_{\text{meas}}}{\sigma_{\text{ref}}} \quad (2)$$

and

$$\beta_{0,LR} = \bar{u}_{\text{meas}} - \beta_{1,LR} \cdot \bar{u}_{\text{ref}}, \quad (3)$$

where σ_{meas} and σ_{ref} represent the standard deviation of reference and measurement data in the measurement period, and $r_{\text{ref,meas}}$ the Pearson correlation coefficient of the respective data. The bar denotes the mean. In the correction period, the relationship is applied to each of the time-series values of the reference data U_{ref} yielding the corrected wind speed values U_{corr} :

$$U_{\text{corr}} = \beta_{0,LR} + \beta_{1,LR} \cdot U_{\text{ref}}. \quad (4)$$

A disadvantage of this model is that the variance of the corrected data u_{corr} is reduced in comparison to the measured data u_{meas} :

$$\begin{aligned} \text{Var}(u_{\text{corr}}) &= \beta_{1,LR}^2 \cdot \text{Var}(u_{\text{ref}}) \\ &= r_{\text{ref,meas}}^2 \cdot \frac{\sigma_{\text{meas}}^2}{\sigma_{\text{ref}}^2} \cdot \text{Var}(u_{\text{ref}}) \\ &= r_{\text{ref,meas}}^2 \cdot \text{Var}(u_{\text{meas}}) \end{aligned} \quad (5)$$

This yields $\text{Var}(u_{\text{corr}}) < \text{Var}(u_{\text{meas}})$ as, in practical applications, the correlation coefficient $r_{\text{ref,meas}} < 1$. Therefore, simple linear regression can be considered a method which generally yields accurate mean wind speeds (Bass et al., 2000; Rogers et al., 2005a; Romo Perea et al., 2011; Weekes and Tomlin, 2014a; Jie Zhang et al., 2014) but not accurate variances; hence, biased estimates of wind speed distribution and energy production can be expected.

A model which addresses this shortcoming and further develops the simple linear regression approach is the *Linear Regression With Residuals* (LR) method discussed in Weekes and Tomlin (2014a). In contrast to simple linear regression, the residuals are explicitly considered, giving the missing variance to the corrected data:

$$U_{\text{corr}} = \beta_{0,LR} + \beta_{1,LR} \cdot U_{\text{ref}} + \varepsilon_{\text{rand}}. \quad (6)$$

$\varepsilon_{\text{rand}}$ is randomly drawn from a normal distribution $\varepsilon_{\text{rand}} \sim \mathcal{N}(\mu, \sigma_{\varepsilon})$ with mean μ and standard deviation σ_{ε} . μ is set to $\mu = 0$ so that the mean value of the corrected wind speeds U_{corr} is not changed. The parameter σ_{ε} can be estimated using the data from the measurement period (Weekes and Tomlin, 2014a). In this context, the deviations of the data points from the regression line (applying simple linear regression) are determined; their standard deviation then yields σ_{ε} . Hence, the induced scatter resembles the scatter which is observed in the measurement period. Weekes and Tomlin (2014a) show that the LR method yields precise mean wind speeds as well as accurate mean wind power densities.

In Rogers et al. (2005a), the *Variance Ratio* (VR) method is proposed as an alternative to the classical linear regression methods. This approach is closely related to (simple) linear regression; in contrast, however, the regression parameters $\beta_{0,VR}$ and $\beta_{1,VR}$ are not calculated using the linear least square method. Instead, $\beta_{1,VR}$ is defined as

$$\beta_{1,VR} = \frac{\sigma_{\text{meas}}}{\sigma_{\text{ref}}}. \quad (7)$$

205 which resembles the particular case of a simple linear regression with correlation coefficient $r_{\text{ref,meas}} = 1$ (compare Eq. (2)). This choice of $\beta_{1,VR}$ ensures that the variance is maintained, in terms of equal variances of measured data u_{meas} and corrected data u_{corr} in the measurement period. $\beta_{0,VR}$ is then computed using Eq. (3) accordingly. This, in turn, ensures that the mean values of measured and corrected data (in the measurement period) are equal. The VR approach therefore maintains both the first and the second order statistical moment of the measured time series in the LTC. Correction is performed via Eq. (4) using
210 the respective regression parameters $\beta_{0,VR}$ and $\beta_{1,VR}$.

In Rogers et al. (2005a) the authors found that the VR method yielded accurate predictions of all investigated metrics including mean wind speed and wind speed distribution. Other studies confirm the suitability of the VR method in the context of long-term correction of wind measurements (see e.g., Weekes and Tomlin, 2014a; Weekes et al., 2015).

3.3 Statistical analysis and definition of error scores

215 For each MCP calculation according to Sect. 3.1, a one-year time series is generated. Based on comparison with the measured one-year data, the following error scores are derived to evaluate the accuracy of these time series:

1. Bias in (annual) mean wind speed, $Err_{\text{mean}} = \frac{\bar{u}_{\text{corr}} - \bar{u}_{\text{meas}}}{\bar{u}_{\text{meas}}}$ (where the bar denotes the respective one-year mean wind speeds).
2. Bias in variance of the (one-year) time series, $Err_{\text{var}} = \frac{Var(u_{\text{corr}}) - Var(u_{\text{meas}})}{Var(u_{\text{meas}})}$.
- 220 3. Bias in energy density, Err_{ED}

As relative values are addressed only, the bias in energy density is simply based on the bias in cubed wind speed u^3 here. The exact procedure of calculation is given in the text of the respective section (Sect. 5.4.3).

4. Bias in theoretical annual energy production of a wind turbine, Err_{turbine}

225 To derive this error score, the theoretical one-year energy production of a wind turbine is calculated using the power curve of a 3.2 MW wind turbine (see Enercon, 2019). This power curve has a cut-in wind speed at 2 m/s and the nominal power is reached at wind speeds of 14 m/s. When the winds are stronger than 25 m/s, no energy is converted (cut-out wind speed). Err_{turbine} is given by the relative deviation of the energy values calculated from the corrected and the measured one-year time series (i.e., similar to Err_{mean} and Err_{var}). Two further power curves with significantly lower

and higher cut-in and cut-out wind speeds (nominal power: 1.8 MW and 4.2 MW) were used in order to quantify the variability for different power curves. As the results only differed slightly and the essential conclusions remained the same, only the results for this 3.2 MW turbine power curve are presented in this study.

4 Theoretical considerations

Before experimental analysis is presented, in this section theoretical aspects are discussed. It should be noted that these theoretical considerations are, to some extent, also valid for a long-term assessment which is based on an entire year of measurement data (i.e., as most commonly done in wind resource assessment today). In this case, the inter-annual variations of the wind conditions represent the key factor. However, these are usually smaller than the seasonal variations during the year, which are discussed below.

4.1 Influence of mean and variance on the estimate of energy

Both mean and variance of the predicted wind speed distribution have an impact on the estimate of the power production of a wind turbine which is, eventually, the target value of a wind resource assessment when planning a wind park. In this section, the importance of an error in each of the two statistical metrics is investigated.

For this purpose, the power in wind P is analyzed. It is known that P scales with the wind speed in third power (u^3). Hence, the expected value $E[P]$ of a wind speed distribution is mainly characterized by (is proportional to) $E[u^3]$. Romo Perea et al. (2011) give an approximation for $E[u^3]$ based on the first three statistical moments of the wind speed distribution,

$$E[u^3] = \bar{u}^3 + 3 \cdot \bar{u} \cdot \sigma_u^2 + \gamma \cdot \sigma_u^3, \quad (8)$$

with σ_u representing the sample standard deviation of wind speeds u and γ the skewness coefficient. The bar denotes the mean. Generally, γ is rather small (Romo Perea et al., 2011) and the term $\gamma \cdot \sigma_u^3$ therefore will be neglected in the following.

Applying the (simplified) formula of the Taylor series method for propagation of error (see e.g., Coleman, 2009),

$$\Delta E[u^3] = \frac{\partial E[u^3]}{\partial \bar{u}} \cdot \Delta \bar{u} + \frac{\partial E[u^3]}{\partial \sigma_u^2} \cdot \Delta \sigma_u^2, \quad (9)$$

with Δ symbolizing the error of the respective parameter, yields

$$\frac{\Delta E[u^3]}{E[u^3]} = \left(1 + \frac{2}{1 + \frac{3}{A}}\right) \cdot \frac{\Delta \bar{u}}{\bar{u}} + \frac{1}{1 + \frac{A}{3}} \cdot \frac{\Delta \sigma_u^2}{\sigma_u^2} \quad (10)$$

as a formula for the overall relative error of $E[u^3]$. The substitution $A = \bar{u}^2 / \sigma_u^2$ was introduced for means of readability.

The available one-year measurement data (see Sect. 2) were used to derive values for A which typically occur at the investigated sites. It was found that $A = 5.0 \pm 0.8$ (mean ± 1 standard deviation). Inserting in Eq. (10) shows that the effect of a relative error in mean wind speed is weighted six times as strong as the relative error in variance σ_u^2 .

Note that simplifications were applied (e.g., neglectation of the skewness of the distribution) and that the output of Eq. (10) varies from site to site (due to a site-dependence of the parameter A). However, a clear impression of a much larger importance of a high accuracy in mean than in the variance of the wind speed distribution is obtained. As will be shown in the experimental section (Sect. 5), the errors in variance can be quite large when a long-term correction of short-term wind measurements is performed and, hence, should not be neglected nevertheless.

Following these considerations, the sections below address the question which factors influence the accuracy of the estimation of the mean and the variance when a long-term correction is performed based on one of the two linear MCP approaches.

4.2 Considerations on seasonal bias in mean wind speed

In both cases of the VR and the LR method, the mean value of the corrected wind speed data is given by

$$\bar{U}_{\text{corr}} = \beta_0 + \beta_1 \cdot \bar{U}_{\text{ref}}, \quad (11)$$

with the respective values of regression parameters β_0 and β_1 .

Using the definition of β_0 (see Eq. (3)) leads to

$$\bar{U}_{\text{corr}} = \bar{u}_{\text{meas}} - \beta_1 \cdot (\bar{u}_{\text{ref}} - \bar{U}_{\text{ref}}). \quad (12)$$

The error in mean wind speed is usually defined as the deviation of the calculated mean wind speed from the "true" value. Hence, the difference $\bar{U}_{\text{corr}} - \bar{U}_{\text{meas}}$ provides a convenient formula for the theoretical (absolute) bias in mean wind speed

$$\begin{aligned} \text{Err}_{\text{mean,theo}} &= \bar{U}_{\text{corr}} - \bar{U}_{\text{meas}} \\ &= (\bar{u}_{\text{meas}} - \bar{U}_{\text{meas}}) - \beta_1 \cdot (\bar{u}_{\text{ref}} - \bar{U}_{\text{ref}}). \end{aligned} \quad (13)$$

This formula is valid for both the LR and VR method (with respective regression parameter $\beta_{1,\text{LR}}$ or $\beta_{1,\text{VR}}$).

Therefore, three factors have a direct impact on the accuracy in mean wind speed when applying either the VR or LR method:

(I) $\bar{u}_{\text{meas}} - \bar{U}_{\text{meas}}$: Deviation of "true" mean wind conditions (measured data) in measurement and long-term period

This part of Eq. (13) denotes the difference of mean wind speeds in measurement and long-term period. Therefore, it can be interpreted as a measure for the representativeness of the period in which the measurement is carried out. In case of periods of lower wind speeds, this quantity is negative ($\bar{u}_{\text{meas}} < \bar{U}_{\text{meas}}$) while positive values occur in case of periods with strong winds ($\bar{u}_{\text{meas}} > \bar{U}_{\text{meas}}$).

(II) $\bar{u}_{\text{ref}} - \bar{U}_{\text{ref}}$: Deviation of the mean wind speeds of the reanalysis data in measurement and long-term period

Similarly to term (I) but related to the reanalysis data, this term reflects the differences of wind conditions in measurement and long-term period given by the reanalysis data.

(III) Regression parameter β_1

285 The regression parameter β_1 weights term (II) and, therefore, determines whether the first or the second part of Eq. (13) dominates. As β_1 is different for the LR and the VR method, the respective results of an LTC will inevitably show differences, accordingly.

Obviously, the value of $Err_{\text{mean,theo}}$ is zero when the terms $\bar{u}_{\text{meas}} - \bar{U}_{\text{meas}}$ and $\beta_1 \cdot (\bar{u}_{\text{ref}} - \bar{U}_{\text{ref}})$ cancel out. While $\bar{u}_{\text{meas}} - \bar{U}_{\text{meas}}$ solely depends on the selected measurement period and the specific site, $\bar{u}_{\text{ref}} - \bar{U}_{\text{ref}}$ is, additionally, highly sensitive to
290 the selected reference data set (reanalysis data in this study) and its capability to reflect the measured annual cycle on site. β_1 , in turn, is dependent on the selected MCP method and can vary in time. In case of representative wind conditions (i.e., small values of terms (I) and (II)), the exact value of β_1 is of minor importance.

4.3 Considerations on seasonal bias in variance

Similarly to the considerations on mean wind speed above, in this section a theoretical perspective on the accuracy in variance
295 is given. For the variance of the corrected data $Var(U_{\text{corr}})$ the following relationship is obtained for the VR method:

$$Var(U_{\text{corr}}) = \beta_{1,\text{VR}}^2 \cdot Var(U_{\text{ref}}) = Var(u_{\text{meas}}) \cdot \frac{Var(U_{\text{ref}})}{Var(u_{\text{ref}})}. \quad (14)$$

The accuracy of the LTC in variance, therefore, directly depends on the representativeness of the measured variance for the long-term period. Furthermore, the ratio of the variances given by the reanalysis data needs to be similar in the correlation and the correction period to yield accurate results. The general accuracy of the reanalysis data regarding the variance, in contrast,
300 is of minor importance.

When the LR method is applied, the respective formula reads:

$$Var(U_{\text{corr}}) = r_{\text{ref,meas}}^2 \cdot Var(u_{\text{meas}}) \cdot \frac{Var(U_{\text{ref}})}{Var(u_{\text{ref}})} + Var(\varepsilon_{\text{rand}}). \quad (15)$$

Hence, the variance of the output data is mainly influenced by three factors here:

- 305
1. the accuracy of the reanalysis data in reproducing the annual variability of variance (similarly as discussed for the VR method)
 2. the correlation coefficient (in the context of $\beta_{1,\text{LR}}$, cf. Eq. (2))

3. the residuals determined in the measurement period or, more specifically, the representativeness of their measured standard deviation $\sigma_\varepsilon = \sqrt{\text{Var}(\varepsilon_{\text{rand}})}$ for the entire correction period (see Sect. 3.2.1)

310 It should be noted that, from a mathematical point of view, factors 2. and 3. are strongly connected (e.g., a lower correlation coefficient implies higher scatter around the linear fit and, hence, variance of the residuals). Therefore, in the experimental section, the analysis is focused on factors 1. and 2. (note that the correlation coefficient contributes as a quadratic term to Eq. (15)).

5 Experimental Results

315 In the following sections, the theoretically derived aspects are further explored and tested experimentally. Afterwards, MCP calculations are presented. Systematic biases are described and discussed. In a last section, the variation of the results between the different sites is explicitly considered.

5.1 Seasonal cycle of mean wind speed in measurement and reanalysis data

Equation (13) in Sect. 4.2 constitutes the essential basis for the understanding of seasonal biases in mean wind speed in the context of long-term correction of wind measurements. According to that formula, both the seasonal cycle of measured wind speed as well as the capability of the reanalysis data to reproduce this course are decisive.

In Central Europe –the region under investigation in this paper– the wind conditions usually show lower mean wind speeds in summer and stronger winds in winter periods (Pryor et al., 2006). The exact seasonal pattern will be different from site to site, depending on site-related properties (e.g., proximity to sea or topographical conditions). In Fig. 2 the average seasonal cycle at the 18 sites as given by the different reanalysis data sets is presented. Additionally, the measured seasonal cycle is shown (black dashed line). In all cases, relative values were used, i.e., the mean wind speeds in the different 90-day periods (see Sect. 3.1) were divided by the annual means of the respective data sets.

All data confirm the typical seasonal pattern described above. Hence, both terms **(I)** and **(II)** in Eq. (13) (i.e., the deviations of the mean wind speeds in short-term and long-term period in measurement or reanalysis data, respectively) will be negative in summer and positive in winter.

For all reanalysis data sets, however, the seasonal variations are over-pronounced in comparison to the measured ones. In the transitional seasons (spring, fall), the deviations of (relative) reanalysis and measured wind speeds are smallest on average. The amplitudes of the curves in Fig. 2 differ, indicating clear differences between the reanalysis data sets.

In order to further analyze this aspect, a parameter d_{mean} was calculated aiming to display the deviations from reanalysis to measured data in the seasonal course. d_{mean} is derived based on mean values of reanalysis (\bar{u}_{ref}) and measurement data (\bar{u}_{meas}) during the 90-day periods in relation to their overall annual mean values (\bar{U}_{ref} and \bar{U}_{meas} , respectively):

$$d_{\text{mean}} = \frac{\bar{u}_{\text{ref}}}{\bar{U}_{\text{ref}}} - \frac{\bar{u}_{\text{meas}}}{\bar{U}_{\text{meas}}}. \quad (16)$$

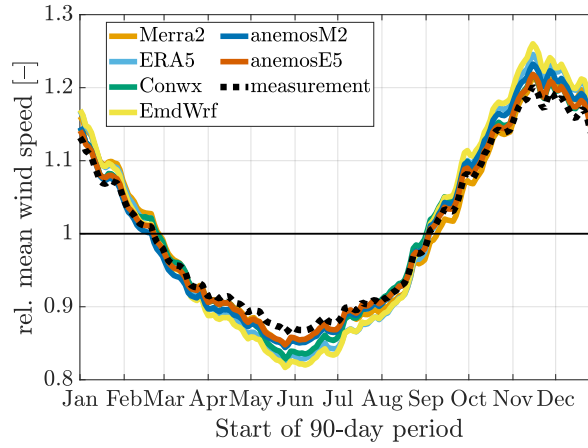


Figure 2. Average annual course of (normalized) wind speed in reanalysis and measurement data. Normalization was done by dividing the mean wind speeds observed in the 90-day periods by the respective annual mean. The individual results obtained at the 18 sites were then averaged arithmetically.

This quantity, therefore, represents the difference between the colored lines and the measured seasonal course (black line) in Fig. 2. For each short-term period, one value of d_{mean} per site and reanalysis data set is derived. Afterwards, values averaged over all sites are calculated resulting in one set of d_{mean} values for each reanalysis data set.

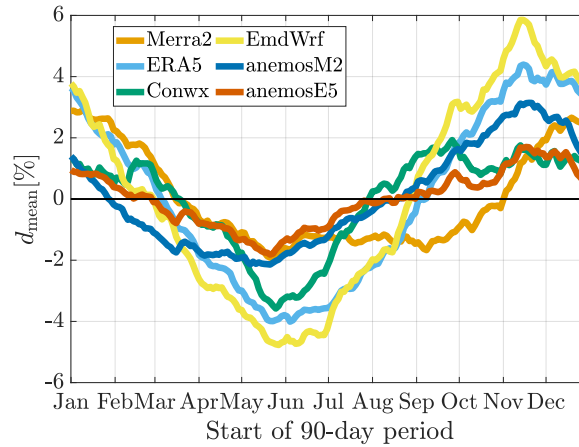


Figure 3. Deviation between reanalysis and measurement data in (normalized) mean wind speed (period of 90 days, arithmetically averaged over all sites).

Figure 3 shows the seasonal course of d_{mean} . Relatively large differences among the different reanalysis data sets can be observed. The aforementioned over-pronounced seasonal course leads to negative deviations in summer and positive values in

winter periods for all reanalysis data sets. Comparing the global reanalysis data sets MERRA-2 and ERA5 with respect to the accuracy in seasonal course shows advantages for the "older" MERRA-2 data set, as a lower amplitude in Fig. 3 is present. This holds true despite or *because of* the fact that the MERRA-2 data are provided at lower heights (50 m, see Sect. 2). This could generally be expected to yield in a lower representativeness regarding the seasonal course at the measurement height. However, the ERA5-based anemosE5 data give better results than the MERRA-2 based anemosM2 data. This might be caused by the further developments by anemos when generating the anemosE5 model (e.g., the additional seasonal correction or the remodeling, see Sect. 2).

350 5.2 Seasonal variations of regression parameter β_1 and correlation coefficient $r_{\text{ref, meas}}$

In addition to the aspects regarding the seasonal course of the wind, Eq. (13) underlines that the magnitude of the regression parameter β_1 plays a significant role. Therefore, in this section β_1 and, in particular, its differences with regard to the two MCP methods, are investigated in more detail.

Comparing the respective definitions of β_1 (Eq. (2) and Eq. (7)) shows that, as mentioned above, the VR method always produces larger slopes than the LR method. Fig. 4 (a) and (b) depict average regression parameters $\beta_{1, \text{VR}}$ and $\beta_{1, \text{LR}}$ and their temporal variation during the year. The respective values were calculated during 90-day periods and arithmetically averaged over all sites.

In contrast to $\beta_{1, \text{VR}}$, $\beta_{1, \text{LR}}$ is subject to clear temporal variations showing lower values in summer and higher values in winter. This, again, reflects the influence of the correlation coefficient $r_{\text{ref, meas}}$ which is only considered explicitly in the LR method and which exhibits a seasonal pattern itself. This pattern is depicted in Fig. 5 where normalized values of $r_{\text{ref, meas}}$ are shown (similarly to the β_1 values in Fig. 4, these were averaged arithmetically over all sites). The correlation coefficient shows a clear seasonal variation for all reanalysis data and decreases significantly towards the summer periods. More unstable stratification and generally lower wind speeds (see Sect. 5.1) might be possible reasons.

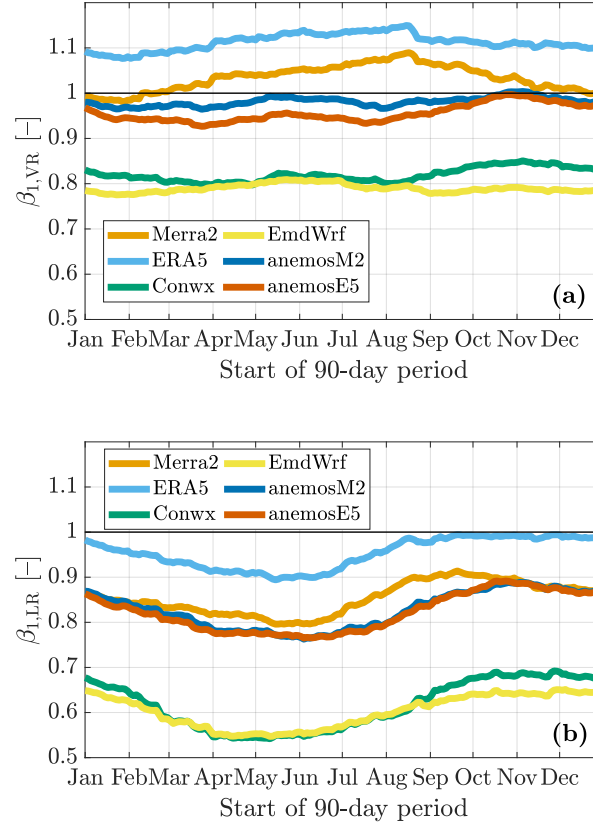


Figure 4. Temporal variation of the regression parameter (a) $\beta_{1,VR}$ for the Variance Ratio and (b) $\beta_{1,LR}$ for the Linear Regression with Residuals method. The respective values were determined using a 90-day sliding window and arithmetically averaged over all sites.

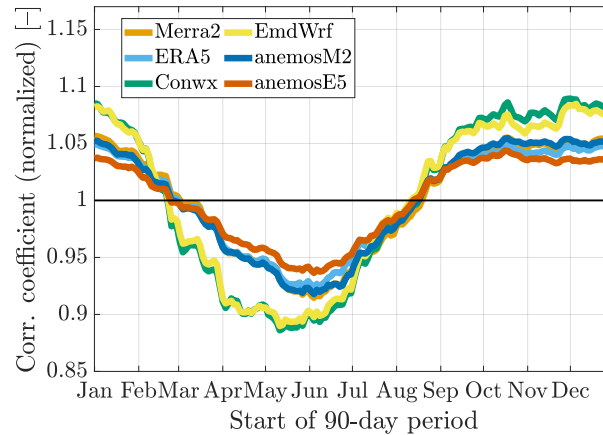


Figure 5. Normalized linear correlation coefficient between measurement and reanalysis data (periods of 90 days, arithmetically averaged over all sites). In the context of normalization the curves were shifted to a mean of 1 to better identify the (relative) temporal variations during the year.

According to Eq. (13), the respective β_1 value weights the seasonal course of the reanalysis data in the determination of the bias in mean wind speed. As a consequence of the findings here, the over-pronounced seasonal cycle of the reanalysis data as depicted above is weighted stronger in winter than in summer periods when the LR approach is applied. Moreover, lower weighting (in comparison to the VR method) occurs throughout as $\beta_{1,VR} > \beta_{1,LR}$.

5.3 Reproduction of the temporal variation of variance in the reanalysis data

As was shown above, the capability of the reanalysis data in reproducing the seasonal course of variance correctly is decisive for an accurate variance of the generated time series. According to the considerations in Sect. 4.3, this is important in case of both MCP methods. As $\beta_{1,VR} = \frac{\sigma_{meas}}{\sigma_{ref}}$ (see Eq. (7)), the seasonal course of the regression parameter $\beta_{1,VR}$ depicted in Fig. 4 (a) gives an impression of how the reanalysis data reproduce the variance and its temporal variation. In order to further investigate this aspect, a measure d_{var} is calculated. Similarly to d_{mean} in Sect. 5.1, d_{var} is defined via the difference of relative values in the 90-day periods,

$$d_{var} = \frac{\text{Var}(u_{ref})}{\text{Var}(U_{ref})} - \frac{\text{Var}(u_{meas})}{\text{Var}(U_{meas})}. \quad (17)$$

Figure 6 shows how the temporal variation of the measured variance throughout the year is reproduced by the different reanalysis data sets.

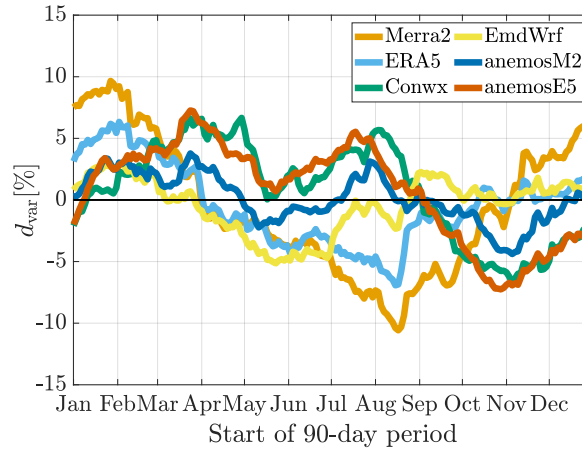


Figure 6. Deviation from reanalysis to measurement data in (normalized) variance (period of 90 days, arithmetically averaged over all sites).

The differences in variance reach values of up to $\pm 10\%$ and are, therefore, generally higher than the deviations in mean wind speed (see Fig. 3). No universal seasonal dependence can be determined as it was observed for the mean wind speed. Some curves in Fig. 6 show minima in summer and high values in winter or spring while others show contrary characteristics.

5.4 MCP calculations: Seasonal bias in mean, variance, and energy

MCP calculations based on 90 days of measurement are now presented. For each reanalysis data set, an average value of the individual error scores related to one measurement period is calculated by arithmetically averaging over all sites. Due to their importance in the theoretical considerations the focus of the analysis is put on mean and variance of wind speed first. Afterwards, seasonal biases in energy density as well as the (theoretical) energy production of a wind turbine are analyzed. In this context, the influence of the systematic biases in both mean and variance on the accuracy in energy is investigated on an experimental level. The analysis in these sections is focused on the systematic biases. The variability of the results (standard deviation) is presented and discussed in a dedicated section afterwards (Sect. 5.4.4).

5.4.1 Seasonal bias in mean wind speed

Figure 7 (a) shows the experimentally obtained bias in mean wind speed (error score Err_{mean}) using the VR method. An inverse shape to the curves of d_{mean} (i.e., the "error" of the reanalysis data in the seasonal course, see Fig. 3) can be observed: A measurement in summer months results in a positive bias in the corrected wind-speed time series while a negative bias is produced when the measurement is conducted in winter. Thus, a positive bias is produced when the reanalysis data underestimate the (relative) mean wind conditions which prevail in the measurement period and vice versa. These findings are valid for all reanalysis data sets although it should be noted that the shapes of the related curves in d_{mean} are not transformed in the (inverse) course of Err_{mean} in exactly the same way.

Strong differences to these observations and even contrary behaviour can be found when the LR method is used (Fig. 7 (b)). For all reanalysis data sets except ERA5, the mean of the corrected wind speed time series is underestimated in case of measurements in summer, while overestimations prevail for winter measurements. The patterns seem not to be directly related to how the reanalysis data reproduce the measured seasonal course of the mean wind speed. Moreover, the ERA5 data gives an inverse curve to all the other reanalysis data sets despite of a high similarity in d_{mean} (Fig. 3). The amplitude of the respective curve is very small indicating a small dependence of the result on the measurement period and, hence, only small seasonal biases.

For most other data sets, the amplitudes of the curves in Fig. 7 (a) and (b) are of comparable magnitude with a slight advantage for the LR method in predicting the mean of the corrected wind-speed time series.

Despite a high similarity in mathematical prospect, the two linear MCP methods yield significantly different results in Err_{mean} . The theoretical analysis of the bias in mean wind speed (Sect. 4.3) yielded a theoretical dependence of Err_{mean} on 1.) the representativeness of the measurement period for the long-term wind conditions, 2.) connected to that, the similarity of the seasonal course in reanalysis data to the measured one, and 3.) the regression parameter β_1 . As 1.) and 2.) are similar for each reanalysis data set, the differences of the results in Fig. 7 (a) and (b), therefore, must have their reasons in differences in β_1 .

As stated above, the VR method provides larger values here than the LR approach (see Sect. 5.2). This leads to the fact that, generally, the seasonal course of the reanalysis data (term $\bar{u}_{\text{ref}} - \bar{U}_{\text{ref}}$ in Eq. (13)) is weighted stronger when the VR method is

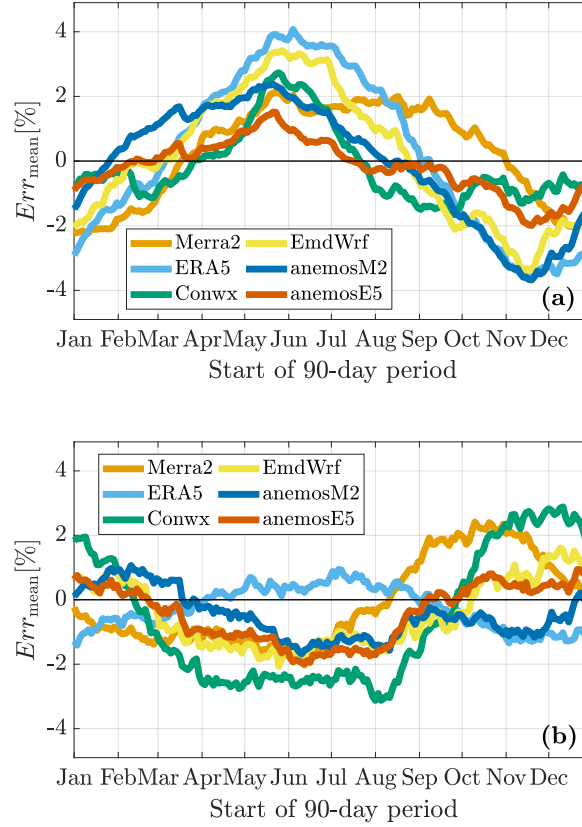


Figure 7. Temporal variation during the year of the bias in mean wind speed using the (a) Variance Ratio, (b) Linear Regression with Residuals method.

used. As a consequence, the effect of the over-pronounced seasonal course of the reanalysis data (see Fig. 2 and 3) dominates here. This is underlined by the fact that Err_{mean} and d_{mean} roughly show inverse shapes. For the LR approach, in contrast, the seasonal course of the reanalysis data is weighted less due to smaller $\beta_{1,\text{LR}}$ values. Therefore, in most instances the seasonal pattern measured on site (term $\bar{u}_{\text{meas}} - \bar{U}_{\text{meas}}$ in Eq. (13)) dominates the overall result. Consequently, most curves of Err_{mean} show a high degree of similarity to the patterns observed in Fig. 2.

As was shown in Fig. 4, in case of the ERA5 data relatively high $\beta_{1,\text{LR}}$ values were obtained. For the LR method this causes a balancing effect (even slightly "overbalanced"). Thus, a relatively small amplitude of Err_{mean} can be observed in Fig. 7 (b) despite, or rather *because of* the overpronounced annual cycle of the ERA5 data. With regard to the VR method, again, highest slopes ($\beta_{1,\text{VR}}$ values) were observed for ERA5 compared to the other reanalysis data sets. As a direct consequence, the product

of regression parameter β_1 and (over-pronounced) seasonal course in the reanalysis data clearly dominates the result of Eq. (13) and the highest amplitude can be observed in Fig. 7 (a).

425 One further example is analyzed briefly here. The largest deviation in the seasonal course d_{mean} was found for the EMD-Wrf Europe+ data set (see Fig. 3). In contrast to the ERA5 data, though, remarkably lower $\beta_{1,\text{LR}}$ values are present for this reanalysis data set (see Fig. 4 (b)). Eventually, the product of (small) regression parameter and (large) deviation of the reanalysis data in the seasonal course in Eq. (13) results in a relatively small amplitude of Err_{mean} .

In summary, it can be stated that the capability of the reanalysis data in reproducing the seasonal course of the "true" wind
430 conditions on site is an important aspect when considering the bias in mean wind speed. However, positive (or negative) deviations in seasonal course do not transform to negative (or positive) biases directly. The regression parameter, depending on both the MCP method and the selected reanalysis data set, strongly influences the outcome additionally.

Note that the influence of the seasonality in $\beta_{1,\text{LR}}$ as shown in Fig. 4 (b) can not be determined exactly here, as the lower values in summer coincide with a stronger effect of the over-pronounced seasonal cycle of the reanalysis data (lower d_{mean}
435 values).

In a study of Bass et al. (2000), long-term measurements instead of reanalyses were used as reference data. 41 pairs of site and reference data in Europe and the US with different terrain types were deployed to test a variety of MCP methods including linear models like linear regression as well as a neural network approach. Hence, long-term corrections of one-year on-site data were performed. Regarding the bias in mean wind speed they found that none of the investigated methods stood out in
440 comparison to the others. It was concluded that the success of the methods "is less to do with the mechanics of the methodology itself, and more to do with facets of the data being analysed". Carta et al. (2013) confirms that the uncertainty of the long-term predictions depends much more on the (reference) data than on the MCP method.

With regard to an LTC of short-term wind measurements, the results of this work only partly agree with these findings from literature. It was shown both theoretically and experimentally that, concerning systematic, seasonal biases, a strong dependence
445 on the selected MCP method occurs. The results above show that very different outcomes can be observed when relatively similar, linear MCP approaches are applied even when the same reference data set is used. In this context, it should be noted that when a long-term correction of entire one-year measurement data is performed, the seasonal aspects discussed here can be replaced by inter-annual variations (which are, however, much smaller); Eq. (13) retains its validity in this case.

In a study of Weekes and Tomlin (2014a) seasonal patterns in the long-term correction of short-term wind measurements
450 are addressed briefly. For both LR and VR, larger biases in mean wind speed were observed when measuring in summer while smaller (more negative) values were obtained for winter measurements. The VR method yielded a smaller amplitude and, in contrast to the LR approach, resulted in negative biases throughout. Furthermore, it was concluded that the sign of the bias varied depending on the specific site when the VR method was applied.

Weekes and Tomlin (2014a) related these seasonal effects to temporal changes in synoptic weather patterns and, connected
455 to that, seasonal patterns in wind direction. It has to be noted that Weekes and Tomlin (2014a) used measurements instead of reanalysis data as reference and all data were collected at heights of around 10 to 20 m. The theoretical background derived in Sect. 4.2 is, however, independent of height and origin of the wind data and can be seen valid universally and applicable also

under these conditions. Against this background, it is likely that not all the reference data used in Weekes and Tomlin (2014a) exhibited an over-pronounced seasonal cycle as present for the reanalysis data used in the study here.

460 Saarnak et al. (2014) applied a linear regression approach to wind data from a site on a Swedish island using MERRA reanalysis (predecessor of MERRA-2). Systematic underestimations in a long-term correction were found when short-term data of three-months winter periods were used. Summer measurements, in turn, resulted in positive biases in mean wind speed. Hence, results similar to the ERA5 curve in Fig. 7 were obtained. Explanations for this seasonality were not given in the study.

5.4.2 Seasonal bias in variance

465 In this section, the bias of the MCP predictions with respect to variance is analyzed. Fig. 8 (a) and (b) show the respective error score Err_{var} .

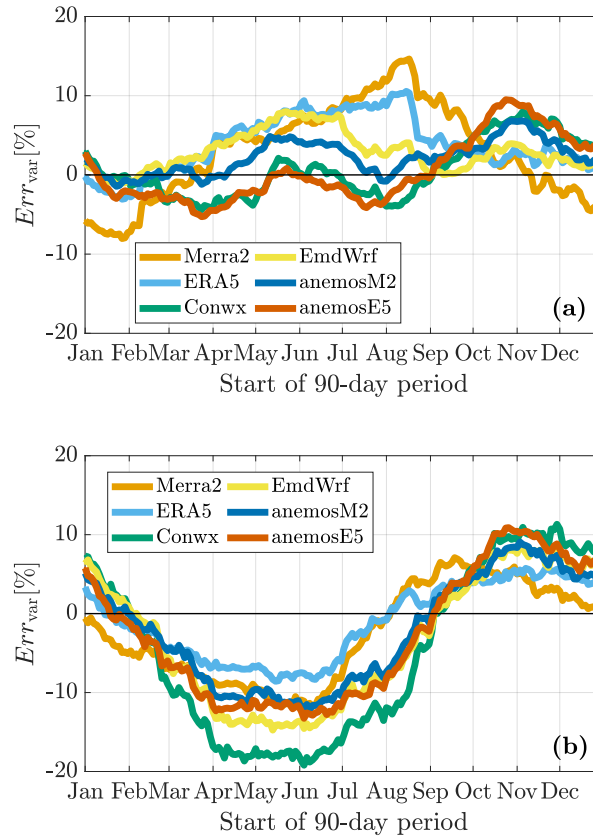


Figure 8. Temporal variation during the year of the bias in variance Err_{var} using (a) Variance Ratio, (b) Linear Regression with Residuals method.

The curves displayed in Fig. 8 (a) for the VR method resemble the inverse course of that observed in Fig. 6, thus, the patterns in the differences in variance. This is not surprising, as the ratio of variances of measurement and reference data is used as a regression parameter in the VR method. Therefore, an error in the temporal variation of the variance given by the reanalysis data has a strong impact on Err_{var} . In summary, the theoretical analysis presented in Sect. 4.3 is confirmed by these experimental results. Connected to that, no clear mean seasonal course can be observed when the VR method is used. The amplitudes of the variations, however, are of distinct magnitude and remarkable errors can be observed.

As shown in Fig. 8 (b) a clear seasonal cycle of Err_{var} is obtained when the LR method is applied. Lower values are present when measuring in summer and higher values can be found in case of winter measurements. This effect can be observed for all reanalysis data sets. In Sect. 4.3 three parameters were identified which have a notable impact on Err_{var} using the LR method. It can be expected that the most important factor is the correlation coefficient $r_{ref, meas}$ as this parameter exhibits a strong seasonal cycle (see Sect. 5.2). Moreover, $r_{ref, meas}$ contributes as a quadratic term to the theoretical calculation of Err_{var} (Eq. (15)). Furthermore, this explains the substantial differences between Fig. 8 (a) and (b), i.e., between the results of the VR and the LR method.

In summary, the amplitudes in Fig. 8 (b) are generally of slightly larger magnitude than those of the variations produced by the VR method. This indicates that the VR method enables to obtain a more accurate variance of the corrected data on average. Differences occur regarding the type of reanalysis data. Similar to the bias in mean wind speed, ERA5 gives the lowest bias in variance when the LR method is used while large biases are obtained when the VR method is applied on the ERA5 data.

5.4.3 Seasonal bias in energy

In Sect. 4.1 a much higher importance of an accuracy in mean than in variance was obtained when aiming for a precise estimate of the energy in the wind. This contrasts with the finding of significantly higher biases in variance than in mean wind speed. In this section, the bias both in energy density (Err_{ED}) as well as in the theoretical energy production of a wind turbine ($Err_{turbine}$) is investigated based on experimental analysis. Emphasis is put on the eventual overall influence of Err_{mean} and Err_{var} , respectively.

Figure 9 (a) and (b) show the error score Err_{ED} . The respective values were obtained according to Eq. (10) using the experimentally derived error values of Err_{mean} and Err_{var} as presented in Sect. 5.4.1 and 5.4.2, respectively. For A the experimentally obtained average value of $A = 5.0$ was used (see Sect. 4.1). Hence, the diagram was produced by a weighted sum of Err_{mean} and Err_{var} .

Additionally, the biases in u^3 based on the time-series values were evaluated experimentally (not shown here). These gave very similar results to that presented in Fig. 9 (a) and (b) indicating that Eq. (8) contains a high validity despite the applied simplifications.

Comparison of Fig. 9 with the plots of Err_{mean} and Err_{var} (Fig. 7 and 8) reveals the influence of the biases in variance and mean wind speed on the bias in energy (density or production). Periods of contrary behaviour of Err_{mean} and Err_{var} (e.g., opposite sign or individual peaks) are most suitable to analyze this aspect here.

500 Generally, the influence of the bias in mean dominates (compare the considerations presented in Sect. 4.1). In some cases, however, the influence of the bias in variance is visible. E.g., in Fig. 9 (b) the sky-blue curve associated to the ERA5 data gives negative values in summer although the related Err_{mean} curve remains positive in this period. This can be traced to the strongly negative Err_{var} values here. When the VR method is used (Fig. 9 (a)), the effect of erroneous variance is even more clearly visible, due to different courses of the respective Err_{mean} and Err_{var} values, e.g., in case of the EMD-ConWx or the
505 anemosE5 data.

The bias in the theoretical energy production of a wind turbine $Err_{turbine}$ is shown directly below in Fig. 9 (c) and (d) allowing a good comparison of the two error scores. The courses of $Err_{turbine}$ show striking differences to the curves of Err_{ED} indicating a large influence of the power curve on the respective results. Comparison with the patterns of Err_{mean} and Err_{var} reveals that the error in mean wind speed is even more decisive for the error in energy production than the theoretical
510 considerations suggest. The seasonal courses in $Err_{turbine}$ are very similar to the seasonal biases in mean wind speed Err_{mean} (see Fig. 7). Its values are approximately twice the ones for the bias in mean wind speed. The influence of the bias in variance obviously is decreased by the power curve and barely visible. This is caused by the effect that variations of very large wind speed values exceeding the rated wind speed of the turbine contribute strongly to variance but do not affect the energy output. However, in specific periods when Err_{var} is large and its seasonal course does not follow the pattern of Err_{mean} , the influence
515 of Err_{var} can be seen. Again, this is most clearly visible in case of the VR method (see, e.g., the data points related to the MERRA-2 or ERA5 data in the mid of August or to the anemosE5 data in fall in Fig. 9 c) in comparison to Fig. 7 a)).

A further difference between Err_{ED} and $Err_{turbine}$ stands out when the VR method is applied. Some curves in Fig. 9 (c) mostly lie above or below zero for the entire year. Such "overall biases" are present especially in the case of the EMD-ConWx (positive overall bias) and the MERRA-2 data (negative overall bias). When applying the LR method (Fig. 9 (d)), hardly any
520 overall bias can be found.

Towards an explanation approach for these overall biases it should be noted that, again, the VR method produces higher values for the slope (β_1) than the LR approach. For the offset (β_0), the same formula is used in both MCP methods, relating offset to slope (see Eq. (3)). As a direct consequence, lower values for the offset are obtained when the VR method is applied. For the VR method, hence, smaller wind speed values are generally corrected towards smaller values, while higher values are
525 increased compared to the correction applied in the LR method. This is visualized in the scatter plot in Fig. 10 where distinct differences between the regression lines can be observed. Hence, wind speeds of small or rather large magnitude are corrected differently. Similar correction is performed for wind speeds near the mean (i.e., values close to 1 in Fig. 10).

This aspect can be expected to average out when considering mean wind speeds. However, it apparently becomes important in case of energy production estimation where the cubic dependence on wind speed as well as the shape of the power curve lead
530 to a different importance (or weighting) of different wind speed values. Eventual wind-speed dependent errors of the reanalysis data can further contribute to this issue.

The reasons for the overall biases, therefore, show to be connected to both, characteristics of the MCP method and the reanalysis data set. Again, the combination of both these facets prove to be decisive with regard to the accuracy of an LTC.

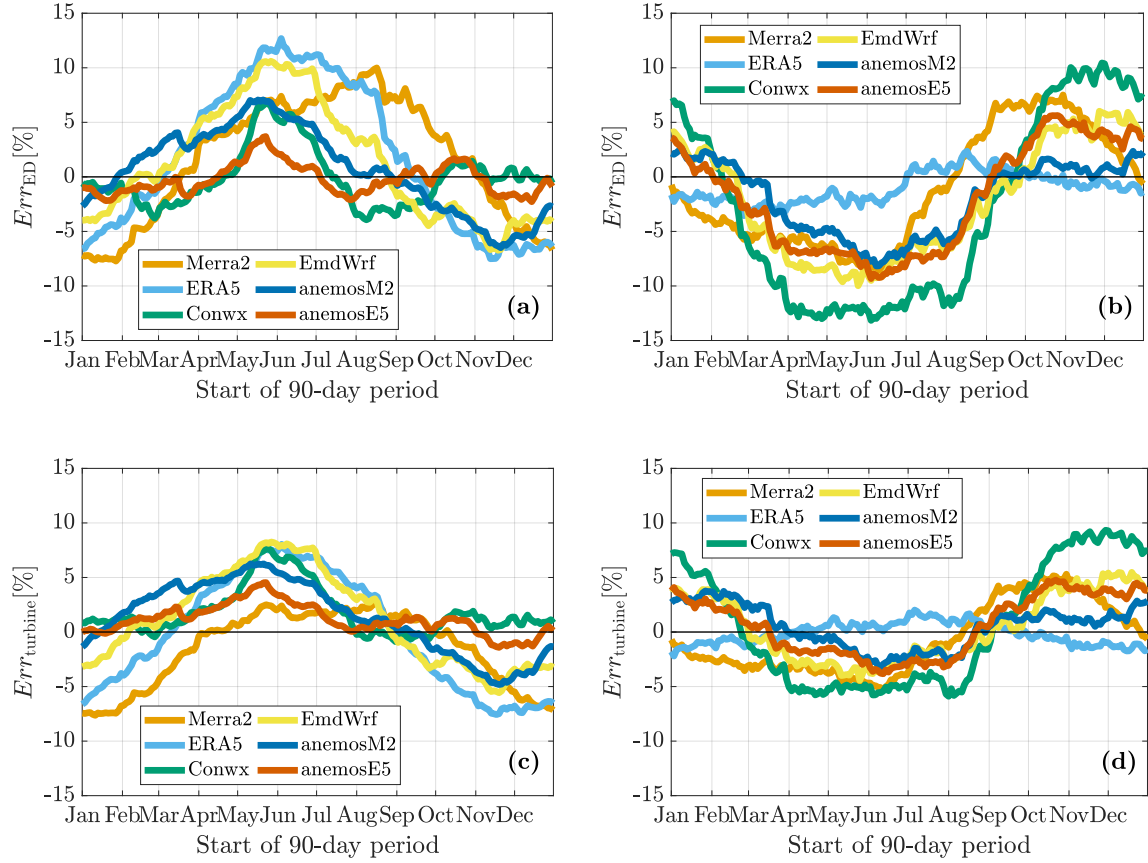


Figure 9. Temporal variation during the year of the bias in the prediction of the energy density Err_{ED} ((a) and (b)) and the theoretical annual energy production of a wind turbine $Err_{turbine}$ ((c) and (d)). The figures on the left ((a) and (c)) refer to the VR method while in figures (b) and (d) the results produced by the LR method are shown.

5.4.4 Variation between the sites

535 Bias values of mean, variance or energy production should not be regarded as the only key figure to describe the accuracy of a
long-term correction procedure as it does not represent the overall uncertainty. In addition, the scatter, i.e., standard deviation
of the individual biases (in terms of variation between the sites) can be judged an important measure as it characterizes the
reliability of the results. Therefore, the standard deviations of Err_{mean} and $Err_{turbine}$ in dependence of the measurement
period are addressed here and shown in Fig. 11. The analysis is restricted to Err_{mean} and $Err_{turbine}$ as these parameters are
540 expected most useful for the wind industry.

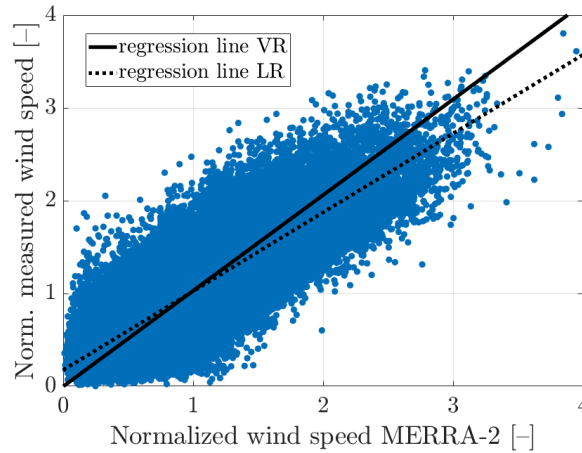


Figure 10. Scatter plot of normalized measured and MERRA-2 data and regression lines to these data using either the VR or the LR method. Normalization was performed by dividing all wind speed values by the overall measured mean. The diagram was produced using the entire measurement data of the 18 sites and the related MERRA-2 data.

Similar to the biases, the variations (standard deviations) are significantly higher for $Err_{turbine}$ than for Err_{mean} . In general, both methods (VR and LR) produce comparable magnitudes while the results, again, strongly depend on the selected reanalysis data. The maximum values for individual reanalysis data sets in Fig. 11 are lowest for the anemos data sets and range from approximately 1 to 5 % in case of Err_{mean} . Differences regarding the MCP method occur in winter periods when considering $Err_{turbine}$ (9 % in maximum values for the VR method and more than 11 % for the LR method). In summary, the variation between the sites is roughly of the same magnitude as the bias values themselves (see Fig. 9).

On average, smallest values can be observed in the beginning of the year and in fall (i.e., measurement period starting in January/February or September/October) for both Err_{mean} and $Err_{turbine}$. This indicates that not only strong biases are present when the measurement is conducted in summer or winter but also higher variations, hence, smaller reliability of these biases can be expected. Once more, this underlines the significance and importance of a sorrow selection of the measurement period, with transitional seasons (spring, fall) to be recommended in Central Europe.

6 Conclusions and outlook

This study delivered in-depth analysis of seasonal effects in the long-term correction of short-term wind measurements. The provided findings can contribute to a further development of reanalysis data as well as improved MCP methods in this respect. In a first step, the importance of the accuracy in mean and variance of wind speed was analyzed with regard to a precise estimate of the energy in the wind. It was shown on a theoretical level, that the relative error in mean contributes six times as much as the relative error in variance in this context. Experimental analysis, in contrast, showed that much larger biases in variance than in mean prevail when MCP predictions are performed (absolute values of more than 15 % were obtained in

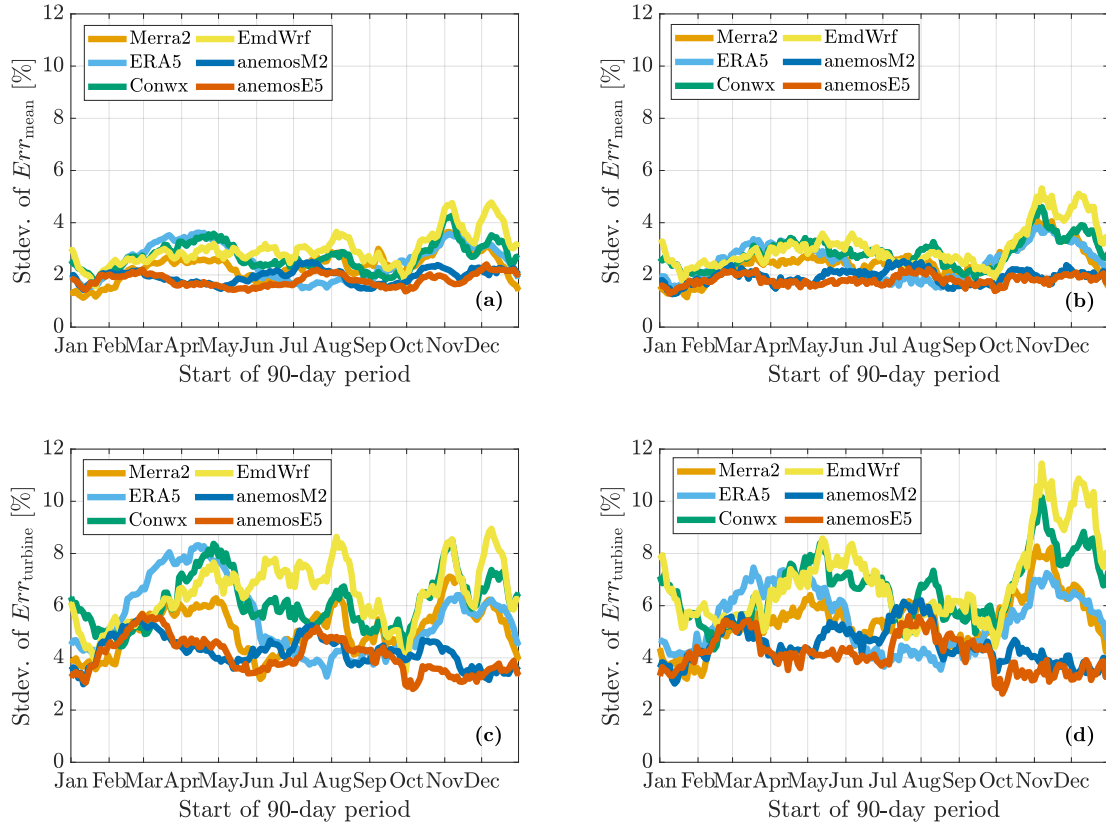


Figure 11. Bias variation between the sites (1 standard deviation) with regard to the accuracy of predicting the mean wind speed ((a) and (b)) and the theoretical energy production of a wind turbine ((c) and (d)). The figures on the left ((a) and (c)) refer to the VR method while in figures (b) and (d) the results produced by the LR method are shown.

comparison to values of $\pm 4\%$, respectively). It was demonstrated that –apart from "overall biases"– the shape of the seasonal course of the bias in mean wind speed was more or less replicated in the bias of the theoretical energy production. Therefore, it can be concluded that a precise estimate of the mean is much important than the correct estimate of the variance when assessing the energy production of a wind turbine.

A formula was derived which delivered the explanation for the seasonal biases in mean wind speed when applying either the Variance Ratio or Linear Regression with Residuals method. It was shown that the representativeness of the measurement period, i.e., the similarity of the wind conditions in correlation and correction period, is important. Moreover, the capability of the reference data to reproduce the seasonal course is significant. Lastly, the regression parameter β_1 (computed differently

for the two MCP methods used in this study) showed to be decisive for the magnitude of the seasonal biases. This theoretical framework was used for explaining the observations.

The largest biases were obtained in case of measurement periods with non-representative wind conditions (i.e., significantly lower or higher mean wind speeds compared to the annual mean – usually summer and winter periods in Central Europe). The magnitude showed to depend on the reference data set. Furthermore, a strong dependence on the MCP method was identified; very different, partly even contrary characteristics in the seasonal biases were found for the VR and LR methods. In contrast to findings of existing publications, hence, this study showed that the biases obtained from a long-term correction of short-term wind measurements (a few months) are connected to both, characteristics of the reference data set as well as the selected MCP method.

In general, measurement periods in transitional seasons (spring, fall) not only resulted in smallest biases but also gave smallest variation between the sites, thus, the highest reliability of the results. The amplitudes of seasonal bias and standard deviation of the results obtained at the individual sites were roughly of same magnitude. If short-term wind measurements are used for wind resource assessments, it is, therefore, highly recommended to conduct these measurements in periods which are likely to be characterized by representative wind conditions (with respect to mean wind speed).

Further research is necessary on how the systematic biases and, finally, the uncertainty of the long-term correction of short-term wind measurements can be reduced in an efficient and expedient way. The authors suggest that this could be approached in different ways. On the one hand, a manual correction based on the experiences described above would reduce the biases. However, the reliability (standard deviation) would not change. A statistics-based approach (e.g., averaging the results of different MCP approaches and/or reference data) as well as machine learning approaches (e.g., learning the seasonal effects from other data sets) might result in larger improvements. On the other hand, the shortcomings of the reference (here: reanalysis) data in reproducing the seasonal course could be addressed. Discrepancies regarding temporal changes in synoptic weather patterns or atmospheric stability processes can be named as possible examples for such weaknesses. The inclusion of further meteorological data reflecting these characteristics could form the basis of a physically motivated approach here. The usefulness of removing seasonal biases in e.g., wind profile extrapolation by including additional parameters like relative humidity was demonstrated in Basse et al. (2020). This approach could also be taken here.

Author contributions. AB had the lead in writing the manuscript and developing the theoretical analysis and methodology for this study. AB also performed all data analysis and visualization. LP contributed to the conceptualisation, development of the methodology, and to writing the manuscript. LP, DC, and AG had a supervisory role during the development of the methodology, data analysis, and the writing process. DC was also responsible for the funding acquisition and the project administration. AG performed valuable preliminary work. All authors revised and edited the manuscript.

Competing interests. The authors declare that they have no conflict of interest.

Acknowledgements. The authors would like to express their gratitude to GWU Umwelttechnik GmbH, Notus Energy, NES GmbH, Meteorological Institute of the University of Hamburg, and Karlsruhe Institute of Technology for providing measurement data. Furthermore, the authors thank EMD Deutschland GbR and anemos GmbH for providing mesoscale reanalysis data.

Financial support. This research was funded by the Federal Ministry of Economic Affairs and Energy (Bundesministerium für Wirtschaft und Energie, BMWi) on the basis of a decision by the German Bundestag, Grant No.: 0324159E.

References

- Albrecht, C. and Klesitz, M.: Long Term Correlation of Wind Measurements Using Neural Networks: A New Method for Post-Processing Short-Time Measurement Data, in: Wind Power Asia 2006, 2006.
- anemos: anemos - Gesellschaft für Umweltmeteorologie mbH: anemos Windatlas D-3km.E5, <https://anemos.de/files/windatlanten/Dokumentation-D-3km.ERA5-standortspezifisch-2020-03.pdf>, 2020a.
- anemos: anemos - Gesellschaft für Umweltmeteorologie mbH: anemos Windatlas D-3km.M2, <https://anemos.de/files/windatlanten/Dokumentation-D-3km.M2-standortspezifisch-2019-02.pdf>, 2020b.
- anemos: anemos - Gesellschaft für Umweltmeteorologie mbH: anemos Windatlas (general information), <https://www.anemos.de/en/windatlas.php>, 2020c.
- anemos: Personal contact via telephone and e-mail with M. Schneider from anemos GmbH in January 2021, 2021.
- Bass, J. H., Rebbeck, M., Landberg, L., Cabré, M., and Hunter, A.: An Improved Measure-Correlate-Predict Algorithm for the Prediction of the Long Term Wind Climate in Regions of Complex Environment: Final Report JOR3-CT98-0295, 2000.
- Basse, A., Callies, D., and Groetzner, A.: Ergebnisbericht zum Round Robin Test "Langzeitkorrektur von Kurzzeitwindmessungen", <http://www.uni-kassel.de/eecs/fachgebiete/integrierte-energiesysteme/aktuelles/nachrichten/article/langzeitkorrektur-von-kurzzeitwindmessungen.html>, 2018.
- Basse, A., Pauscher, L., and Callies, D.: Improving Vertical Wind Speed Extrapolation Using Short-Term Lidar Measurements, Remote Sensing, 12, 1091, <https://doi.org/10.3390/rs12071091>, 2020.
- Bilgili, M., Sahin, B., and Yasar, A.: Application of artificial neural networks for the wind speed prediction of target station using reference stations data, Renewable Energy, 32, 2350–2360, <https://doi.org/10.1016/j.renene.2006.12.001>, <http://www.sciencedirect.com/science/article/pii/S0960148106003429>, 2007.
- Bradley, S.: Atmospheric acoustic remote sensing, CRC Press, Boca Raton, Fla., 2008.
- Carta, J. A., Velázquez, S., and Cabrera, P.: A review of measure-correlate-predict (MCP) methods used to estimate long-term wind characteristics at a target site, Renewable and Sustainable Energy Reviews, 27, 362–400, <https://doi.org/10.1016/j.rser.2013.07.004>, 2013.
- CDS: ERA5: Fifth Generation of ECMWF Atmospheric Reanalyses of the Global Climate. Copernicus Climate Change Service Climate Data Store (CDS). ECMWF. Data accessed June-July 2020, <https://cds.climate.copernicus.eu/cdsapp#!/>, 2018.
- Coleman, H. W.: Experimentation, validation, and uncertainty analysis for engineers, John Wiley & Sons, 2009.
- Corotis, R. B.: Stochastic modelling of site wind characteristics. Final report, <https://doi.org/10.2172/7257559>, 1976.
- Draper, N. R. and Smith, H.: Applied regression analysis, Wiley series in probability and statistics Texts and references section, Wiley, New York and Chichester and Weinheim and Brisbane and Singapore and Toronto, third edition edn., <https://doi.org/10.1002/9781118625590>, <http://onlinelibrary.wiley.com/book/10.1002/9781118625590>, 1998.
- Ellison, S. L. R., Farrant, T. J., and Barwick, V.: Practical statistics for the analytical scientist: A bench guide, Royal Society of Chemistry, Cambridge, 2nd ed. edn., 2009.
- EMD: EMD International A/S: EMD-ConWx, http://www2.emd.dk/admin/helpWiki/index.php/EMD-ConWx_Meso_Data_Europe, 2020a.
- EMD: EMD International A/S: EMD-WRF Europe+, <https://www.emd.dk/data-services/mesoscale-time-series/pre-run-time-series/emd-wrf-europe-mesoscale-data-set>, 2020b.
- Emeis, S., Harris, M., and Banta, R. M.: Boundary-layer anemometry by optical remote sensing for wind energy applications, vol. 16, Meteorologische Zeitschrift, <https://doi.org/10.1127/0941-2948/2007/0225>, 2007.

- 640 Enercon: ENERCON Product Portfolio: Overview of Wind Energy Converters - E-115, <https://www.enercon.de/en/downloads/>, 2019.
- FGW e.V.: Fördergesellschaft Windenergie und andere dezentrale Energien (FGW): Technical Guidelines for Wind Turbines: Determination of Wind Potential and Energy Yield (TR6), 2020.
- García-Rojo, R.: Algorithm for the Estimation of the Long-Term Wind Climate at a Meteorological Mast Using a Joint Probabilistic Approach, *Wind Engineering*, Volume 28, 213–224, 2004.
- 645 GMAO: MERRA-2 `tavg1_2d_slv_Nx: 2d,1-Hourly,Time-Averaged,Single-Level,Assimilation,Single-Level Diagnostics V5.12.4`, <https://doi.org/10.5067/VJAFPLIICSIV>, 2015.
- Hersbach, H. et al.: The ERA5 global reanalysis, *Q.J.R. Meteorol. Soc. (Quarterly Journal of the Royal Meteorological Society)*, 146, 1999–2049, <https://doi.org/10.1002/qj.3803>, 2020.
- IEC: International Electrotechnical Commission: IEC 61400-12 ed. 2: Power Performance Measurements of Electricity Producing Wind
- 650 Turbines, 2017.
- Jie Zhang, Souma Chowdhury, Achille Messac, and Bri-Mathias Hodge: A hybrid measure-correlate-predict method for long-term wind condition assessment, *Energy Conversion and Management*, 87, 697–710, <https://doi.org/10.1016/j.enconman.2014.07.057>, <http://www.sciencedirect.com/science/article/pii/S0196890414006888>, 2014.
- Justus, C., Mani, K., and Mikhail, A.: Interannual and Month-to-Month Variations of Wind Speed, *Journal of Applied Meteorology*, 18, 1979.
- 655 Klink, K.: Trends and Interannual Variability of Wind Speed Distributions in Minnesota, *Journal of Climate*, 15, 3311–3317, 2002.
- Lackner, M. A., Rogers, A. L., and Manwell, J. F.: Uncertainty Analysis in MCP-Based Wind Resource Assessment and Energy Production Estimation, *Journal of Wind Engineering and Industrial Aerodynamics*, 130, 257, <https://doi.org/10.1115/1.2931499>, 2008.
- Leleu, K.: Leosphere Windcube User Guide, Version V.1.2 (March 2019), 2019.
- Liléo, S., Berge, E., Undheim, O., Klinkert, R., and Bredesen, R. E.: Long-term correction of wind measurements. State-of-the-art, guidelines
- 660 and future work, *Complexity*, pp. 2–3, 2013.
- López, P., Velo, R., and Maseda, F.: Effect of direction on wind speed estimation in complex terrain using neural networks, *Renewable Energy*, 33, 2266–2272, <https://doi.org/10.1016/j.renene.2007.12.020>, 2008.
- MEASNET: Measuring Network of Wind Energy Institutes: Evaluation of Site-Specific Wind Conditions: Version 2 April 2016, 2016.
- Miguel, J. V. P., Fadigas, E. A., and Sauer, I. L.: The Influence of the Wind Measurement Campaign Duration on a Measure-Correlate-Predict
- 665 (MCP)-Based Wind Resource Assessment, *Energies*, 12, 3606, <https://doi.org/10.3390/en12193606>, 2019.
- NASA: Global Modeling and Assimilation Office: Modern-Era Retrospective analysis for Research and Applications, MERRA Version 2, gmao.gsfc.nasa.gov/reanalysis/MERRA-2, 2019.
- Pauscher, L., Callies, D., Klaas, T., and Foken, T.: Wind observations from a forested hill: Relating turbulence statistics to surface characteristics in hilly and patchy terrain, *Meteorologische Zeitschrift*, 27, 43–57, <https://doi.org/10.1127/metz/2017/0863>, <http://dx.doi.org/10.1127/metz/2017/0863>, 2018.
- 670 Pryor, S. C., Barthelmie, R. J., and Schoof, J. T.: Inter-annual variability of wind indices across Europe, *Wind Energy*, 9, 27–38, <https://doi.org/10.1002/we.178>, 2006.
- Pryor, S. C., Shepherd, T. J., and Barthelmie, R. J.: Interannual variability of wind climates and wind turbine annual energy production, *Wind Energy Science*, 3, 651–665, <https://doi.org/10.5194/wes-3-651-2018>, 2018.
- 675 Ramon, J., Lledó, L., Torralba, V., Soret, A., and Doblas-Reyes, F. J.: What global reanalysis best represents near-surface winds?, *Quarterly Journal of the Royal Meteorological Society*, 145, 3236–3251, <https://doi.org/10.1002/qj.3616>, 2019.

- Rogers, A. L., Rogers, J. W., and Manwell, J. F.: Comparison of the performance of four measure–correlate–predict algorithms, *Journal of Wind Engineering and Industrial Aerodynamics*, 93, 243–264, <https://doi.org/10.1016/j.jweia.2004.12.002>, 2005a.
- Rogers, A. L., Rogers, J. W., and Manwell, J. F.: Uncertainties in Results of Measure-Correlate-Predict Analyses, *European Wind Energy Conference and Exhibition 2006, EWEC 2006*, 3, 2005b.
- Romo Perea, A., Amezcua, J., and Probst, O.: Validation of three new measure-correlate-predict models for the long-term prospection of the wind resource, *Journal of Renewable and Sustainable Energy*, 3, 023 105, <https://doi.org/10.1063/1.3574447>, 2011.
- Saarnak, E., Bergström, H., and Söderberg, S.: Uncertainties Connected to Long-Term Correction of Wind Observations, *Wind Engineering*, 38, 233–248, <https://doi.org/10.1260/0309-524X.38.3.233>, 2014.
- 685 Sørensen, J. D., Sørensen, J. D., and Sørensen, J. N.: Wind energy systems: Optimising design and construction for safe and reliable operation, vol. Number 10 of *Woodhead Publishing Series in Energy*, Woodhead Publishing, Cambridge, England, 2011.
- Taylor, M., Mackiewicz, P., Brower, M. C., and Markus, M.: An Analysis of Wind Resource Uncertainty in Energy Production Estimates, AWS Truewind, 2004.
- Velázquez, S., Carta, J. A., and Matías, J. M.: Comparison between ANNs and linear MCP algorithms in the long-term estimation of the
 690 cost per kWh produced by a wind turbine at a candidate site: A case study in the Canary Islands, *Applied Energy*, 88, 3869–3881, <https://doi.org/10.1016/j.apenergy.2011.05.007>, 2011.
- Weekes, S. M. and Tomlin, A. S.: Data efficient measure-correlate-predict approaches to wind resource assessment for small-scale wind energy, *Renewable Energy*, 63, 162–171, <https://doi.org/10.1016/j.renene.2013.08.033>, 2014a.
- Weekes, S. M. and Tomlin, A. S.: Low-cost wind resource assessment for small-scale turbine installations using site pre-screening and
 695 short-term wind measurements, *IET Renewable Power Generation*, 8, 349–358, <https://doi.org/10.1049/iet-rpg.2013.0152>, 2014b.
- Weekes, S. M. and Tomlin, A. S.: Comparison between the bivariate Weibull probability approach and linear regression for assessment of the long-term wind energy resource using MCP, *Renewable Energy*, 68, 529–539, <https://doi.org/10.1016/j.renene.2014.02.020>, 2014c.
- Weekes, S. M., Tomlin, A. S., Vosper, S. B., Skea, A. K., Gallani, M. L., and Standen, J. J.: Long-term wind resource assessment for small and medium-scale turbines using operational forecast data and measure–correlate–predict, *Renewable Energy*, 81, 760–769,
 700 <https://doi.org/10.1016/j.renene.2015.03.066>, 2015.
- WRF: Weather Research And Forecasting Model, <https://www.mmm.ucar.edu/weather-research-and-forecasting-model>, 2020.